



The Flynn effect for verbal and visuospatial short-term and working memory: A cross-temporal meta-analysis



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ABSTRACT

The Flynn effect has been investigated extensively for IQ, but few attempts have been made to study it in relation to working memory (WM). Based on the findings from a cross-temporal meta-analysis using 1754 independent samples ($n = 139,677$), the Flynn effect was observed across a 43-year period, with changes here expressed in terms of correlations (coefficients) between year of publication and mean memory test scores. Specifically, the Flynn effect was found for forward digit span ($r = 0.12, p < 0.01$) and forward Corsi block span ($r = 0.10, p < 0.01$). Moreover, an anti-Flynn effect was found for backward digit span ($r = -0.06, p < 0.01$) and for backward Corsi block span ($r = -0.17, p < 0.01$). Overall, the results support co-occurrence theories that predict simultaneous secular gains in specialized abilities and declines in g. The causes of the differential trajectories are further discussed.

1. Introduction

The phenomenon called the Flynn Effect (FE) is the increase in cognitive test performance in a population over time, as originally observed by Runquist (1936), who found a systematic increase in scores on the Minnesota College Aptitude test between 1929 and 1933. The FE term was coined by Herrnstein and Murray (1994, p. 307), after James Flynn who demonstrated specifically increases in the Intelligence Quotient (IQ) (Flynn, 1984, 1987, 1990, 2012), also confirmed by a range of researchers in this field (e.g. Pietschnig & Voracek, 2015; Pietschnig, Voracek, & Formann, 2010; Trahan, Stuebing, Hiscock, & Fletcher, 2014).

Apart from the FE for IQ, secular gains for working memory (WM), which can be defined as the immediate and transient retention of information for temporary storage and processing, might be also expected given the strong link between this function and IQ (e.g. Ackerman, Beier, & Boyle, 2005; Cohen & Sandberg, 1977; Colom, Abad, Quiroga, Shih, & Flores-Mendoza, 2008; Colom, Abad, Rebollo, & Chun Shih, 2005; Colom, Flores-Mendoza, Quiroga, & Privado, 2005). Indeed, tests of WM such as forward and backward digit span have often been incorporated in IQ test batteries, including the various versions of the Wechsler Adult Intelligence Scale (WAIS) (Wechsler, 1981, 1987, 2008). Here, a common procedure is the forward digit span test, in

which strings of digits are presented for immediate recall, the lengths of strings increasing until the person consistently makes errors, with the backward digit span variant involving reverse order recall of the digit strings (Woods et al., 2011). As might be predicted, a WM FE has been reported using digit span tasks, albeit with inconsistent findings across studies. For example, Wicherts et al. (2004) compared digit span scores between two Dutch standardization cohorts, one from 1967/1968 ($n = 2100$) and the other from 1998/1999 ($n = 77$). The results revealed the smallest gain was on a digit span subscale (effect size = 0.50; S.D. units) in comparison to the highest gain on a similarity subscale (effect size = 1.48; S.D. units). However, Kanaya and Ceci (2011) also investigated a longitudinal FE on digit span task among 672 school children who were assessed in 1974 and again in 2002, finding no difference in scores. Recently, Gignac (2015) analyzed forward and backward digit span data from 10 normative studies conducted between 1923 and 2008 and showed that there were no significant trends in any of them (respectively, $r_{\text{forward}} = 0.45, p = 0.27$ and $r_{\text{backward}} = -0.57, p = 0.12; k = 9$). However, shortly afterwards, Woodley of Menie and Fernandes (2015) re-analyzed Gignac's (2015) data with the same 10 normative samples but using a weighted regression by sample size method. This result produced modest to large associations ($r_{\text{forward}} = 0.58$ and $r_{\text{backward}} = -0.55; n = 6841$), their analysis was the first study to demonstrate robust and opposing

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temporal trends for these indicators. Accordingly, Woodley of Menie and Fernandes (2015)'s statistical technique was thought superior to Gignac's method in terms of statistical power to detect the effect. However, both studies have limitations due to using only a relatively small number of studies (i.e. $k = 10$). Moreover, although verbal WM as indexed by backward digit span tasks has been widely used for a long time in both clinical and non-clinical settings, there are still very few studies that have looked for FE using verbal WM. Thus, further research on this topic with a more substantial sample size is needed.

In parallel, studies have shown a strong relationship between visuospatial WM and cognitive abilities (e.g. Ashkenazi, Rosenberg-Lee, Metcalfe, Swigart, & Menon, 2013; Colom et al., 2005; Colom et al., 2005; Kytälä & Lehto, 2008; Miyake, Friedman, Rettinger, Shah, & Hegarty, 2001). Despite the fact that non-verbal abilities tend to show a larger FE (Pietschnig & Voracek, 2015; Rönnlund, Carlstedt, Blomstedt, Nilsson, & Weinehall, 2013), to our knowledge, the FE for visuospatial WM has received very little attention and no study has yet been reported. The Corsi block task has long been used as a key measure for visuospatial WM since it was created by Corsi in 1972 (Corsi, 1972; Kessels, van Zandvoort, Postma, Kappelle, & de Haan, 2000). This task consists of an array of nine cubical blocks in which the participant has to remember sequences of blocks "tapped" out by an experimenter. As with digit span, the task starts with short sequences and terminates at higher levels when performance is consistently unsuccessful. The participant is required to repeat the given sequence either identically or in the reversal order (Brunetti, Del Gatto, & Delogu, 2014a; Kessels et al., 2000). Since it has been widely used for a long period of time, this task is a rich resource for conducting meta-analyses to address the gap in our understanding concerning a visuospatial WM FE.

Accordingly, the objectives of the present study were to investigate the FE and the anti-FE for both verbal and visuospatial WM using a Cross-Temporal Meta-Analysis (CTMA). This technique provided a possibly beneficial way to observe the change tendency of a set of specific variables over a period of time and to test possible mediators underlying that change (Donnelly & Twenge, 2017; Zhang, Tan, Wu, Han, & Wang, 2016). In this study, the term short-term memory (STM) was used to indicate working memory functioning in which only temporary storage was required, as measured using forward digit span and forward Corsi block, respectively, for verbal and visuospatial STM. This is contrasted with WM, here defined as involving both storage and processing of information, in this study backward digit span and backward Corsi block span used respectively as measures of verbal and visuospatial WM.

2. Methods

2.1. Literature search

The target articles were identified through 15 scientific Databases from 1910 to 2016; i) Taylor & Francis Online, ii) ScienceDirect, iii) SpringerLink, iv) PsycARTICLES & PsychINFO, v) Wiley Online Library, vi) Sage Journals, vii) Oxford Journals, viii) Cambridge Journals, ix) ProQuest dissertations and theses, x) Karger Online Journals, xi) BMJ, xii) JSTOR, xiii) Frontiers in Psychology, xiv) BioMed Central, and xv) HighWire Journal. The following search terms were used individually and in combination using the Boolean OR function to maximize search sensitivity; "digit span*", "Wechsler's digit span", "forward digit span", "backward digit span", "Corsi block*", "forward Corsi block", and "backward Corsi block" (* is a truncation symbol to represent multiple endings and spellings). No limits were set. The search also identified any data published in journals as well as that found in theses or dissertations (grey literature or unpublished material). Results from all e-databases were combined and duplicates removed subsequently. Despite the fact that we included doctoral dissertations in the current dataset, others sources of grey literature were not systematically searched (e.g. unpublished research reports, submitted or in-press articles).

This may have resulted in some related data not being included in this CTMA.

2.2. Inclusion and exclusion criteria

The meta-analysis included data from studies specifically using the Wechsler-Bellevue scales and its variants for forward/backward digit span and those using the standard versions of the forward/backward Corsi block tasks (see Corsi, 1972; Kessels et al., 2000). Included in the final data-set were data only from studies that reported the mean and/or standard deviation of the raw score for 'longest forward or backward spans' (LFS or LBS) recalled without error. Thus, studies were excluded if they reported only total digit span or Corsi block span scores, the scores for number of items (trials) correct or standardized scores as well as age-scaled scores. Also, studies had to use the same design protocols as the original version of the test (Corsi, 1972) in which the number of blocks increased from a minimum of 2 to a maximum of 9 blocks and the participant was given two attempts at each pattern length. In addition, if studies adopted a test-retest method, only the mean and/or standard deviation scores for pre-test or baseline were recorded and if several articles investigated the same sample or used the same dataset, the statistical parameters (means, standard deviations, standard errors) were treated as a single data point. Studies were excluded if they investigated clinical research participants (with mental or physical disabilities) with the exception of when they reported data for control groups, as such data are appropriate for use in this study. Also, review articles, systematic reviews, meta-analyses, research protocols and case report studies were excluded due to the fact that no targeted statistical parameters are reported in these studies.

In short, there were four main memory tests considered, namely Forward Digit Span (FDS), Backward Digit Span (BDS), Forward Corsi Block (FCB), and Backward Corsi Block (BCB); these were designated in the meta-analysis as representing respectively verbal short-term memory (FDS), visuospatial short-term memory (FCB), verbal working memory (BDS), and visuospatial working memory (BCB).

2.3. Moderator variables

Age was used as a moderator variable, because of the possibility that the mean age for each study would account for variations in the relationship between mean scores for memory tests and years of publication, given evidence for aging affecting STM and WM capacities (Bopp & Verhaeghen, 2005; Hester, Kinsella, & Ong, 2004; Myerson, Emery, White, & Hale, 2003; Pelosi & Blumhardt, 1999; Wingfield, Stine, Lahar, & Aberdeen, 1988). Similarly, gender was also used as a moderator, with previous studies showing gender differences on memory tests, with male participants tending to produce higher scores on spatial tasks but female participants tending to have higher scores on verbal tasks (Andreano & Cahill, 2009; Hill, Laird, & Robinson, 2014; Lowe, Mayfield, & Reynolds, 2003). Thus, the relative strengths for male and female participants on memory tasks may confound the interpretation of the FE on STM and WM tasks.

In addition, prior research has suggested possible negative effect of malnutrition, poorer health, and poverty on cognitive abilities of people in developing countries (Grantham-McGregor et al., 2007), especially for memory development (Khor & Misra, 2012). Furthermore, many studies have shown lower fertility rate among women in developed countries than in developing countries (Bureau, 2016; Nargund, 2006). The FE and anti-FE were also found as partly being caused by fertility patterns in many countries (Kong et al., 2017; Menie, Fernandes, José Figueredo, & Meisenberg, 2015; Sundet, Borren, & Tambs, 2008). Hence the study investigated whether or not this aspect affects memory scores of people between the two types of countries over time. Here, type of country was divided into "developing" versus "developed" according to the terminology recently adopted by the International Statistical Institute (ISI) (2016). A further consideration was the test platform of the

Table 1
Sample characteristics.

| Variables | FDS (ks = 742) (ns = 48,955) | | BDS (ks = 594) (ns = 70,424) | | FCB (ks = 307) (ns = 16,514) | | BCB (ks = 111) (ns = 3784) | |
|---|---------------------------------|------------|---------------------------------|-----------|---------------------------------|------------|-------------------------------|-----------|
| | Mean (SD) | Min – Max | Mean (SD) | Min – Max | Mean (SD) | Min – Max | Mean (SD) | Min – Max |
| Years of publication ^a | 2003.15 (15) | 1973–2016 | 2005.33 (15) | 1975–2016 | 2008.22 (9) | 1982–2016 | 2009.71 (6) | 1989–2016 |
| Age of participant ^b | 36.76 (230.90) | 3.65–92.40 | 41.27 (287.11) | 4–92.50 | 29.31 (162.48) | 3.04–86.65 | 43.06 (149.39) | 4–86.65 |
| Mean score ^b | 6.11 (7.22) | 0.81–9.10 | 4.51 (9.88) | 1.08–7.20 | 5.25 (6.35) | 2.20–7.80 | 4.74 (4.81) | 2.10–7.09 |
| SD score ^b | 1.33 (3.41) | 0.10–3.08 | 1.38 (4.34) | 1.08–7.20 | 1.11 (3.25) | 2.20–7.80 | 1.01 (1.90) | 0.17–2.53 |
| Sex ^b | | | | | | | | |
| Male ^c | 5.77 (9.37) | 3.70–7.80 | 4.30 (5.67) | 2.10–6.90 | 4.86 (5.90) | 3.10–7.00 | 5.02 (3.40) | 2.70–5.70 |
| Female ^d | 6.12 (10.76) | 3.40–8.74 | 4.64 (6.86) | 2.10–7.13 | 4.80 (5.92) | 3.70–7.00 | 4.75 (4.90) | 2.10–5.93 |
| Sample size | 65.98 (191.85) | 1–4100 | 118.56 (352.58) | 5–4251 | 53.79 (80.52) | 9–632 | 34.09 (39.56) | 10–362 |
| Test platform | | | | | | | | |
| Non-computerized (paper-and-pencil or wooden cube) (%) | 676 (91.10) | – | 552 (92.90) | – | 229 (74.60) | – | 99 (89.20) | – |
| Computerized (%) | 66 (8.90) | | 42 (7.10) | | 78 (25.40) | | 12 (10.80) | |
| Type of country | | | | | | | | |
| Developed | 673 (90.70) | – | 529 (89.10) | – | 283 (92.20) | – | 87 (78.40) | – |
| Developing | 69 (9.30) | | 65 (10.90) | | 24 (7.80) | | 24 (21.60) | |

ks = Number of independent samples in which means memory test was available for the analysis, ns = Sum of the sample size for each memory test.

FDS = Forward Digit Span; BDS = Backward Digit Span; FCB = Forward Corsi Block; BCB = Backward Corsi Block.

^a Median and interquartile range (IQR).

^b Sample size weights.

^c Weighted mean with 95% Confidence Interval for male; FDS = 5.77 (5.50–6.04); BDS = 4.30 (4.11–4.50); FCB = 4.86 (4.63–5.10); BCB = 5.02 (4.53–5.51).

^d Weighted mean with 95% Confidence Interval for female; FDS = 6.12 (5.82–6.43); BDS = 4.64 (4.37–4.91); FCB = 4.80 (4.59–5.01); BCB = 4.75 (4.20–5.30).

memory tests used and this was included as a moderator in the current analysis. The test platform could be categorized into computerized and non-computerized versions, as shown in Table 1.

2.4. Statistical analyses

The study used the modified meta-analysis technique, CTMA, to analyze the extent to which the mean scores for memory tests changed over decades based on temporal correlation coefficients, utilizing year of study publication and weighted based on sample size (Twenge, 2000). To estimate the magnitude of change among the mean scores over time, a moderated and weighted least squares multiple regression was applied using Statistica 13 and Statgraphics. Both standardized and unstandardized regression weights were reported. The mean age of participant, percentage of men in each study, test platform (computerized vs. non-computerized), and type of country (developed vs. developing) were inputted into the model as moderators. To indicate the magnitude of change in STM and WM scores over time, a weighted linear regression was calculated for each memory test. The year of publication was used as the predictor with the sample mean FDS, BDS, FCB or BCB scores as the outcome measures (Pietschnig et al., 2010).

In the current analysis, the number of individuals (n) was used to estimate the model degrees of freedom and significances. Essentially, the CTMA method is used to analyze trend scores over time (e.g. Twenge, 2000; Twenge & Campbell, 2001; Twenge & Campbell, 2008; Wongupparaj, Kumari, & Morris, 2015). In contrast to other types of meta-analysis, the primary outcome of interest is not the effect size. The temporal relationship between mean score and year of publication of these scores were collected from target studies and analyzed. In addition, this technique was used to analyze sample mean data, also weighting the analysis by sample size from each study, so taking into account the fact that bigger sample sizes should be weighted more highly and hence accurately (see Lipsey & Wilson, 2001).

In a conventional meta-analysis of effect sizes, the validity of the research findings may be affected by publication bias, namely the file drawer effect (Rothstein, Sutton, & Borenstein, 2006). However, a factor mitigating this effect is that the mean scores from each published study was collected and analyzed, rather than the effect sizes from the primary measures of the study. This related to the target mean score for the digit span or Corsi block measures rarely being associated with the

central hypotheses considered by the journal papers, tending to be reported on an incidental basis as descriptive or Supplementary data (Booth, Sharma, & Leader, 2016; Cooper, 1998; Pietschnig et al., 2010; Schmidt & Hunter, 2015).

Initially, outlier and influential case diagnostics were performed using Cook's distances, DFBETAS, and DFFITS (Viechtbauer & Cheung, 2010). First, according to Skrondal and Rabe-Hesketh (2004), for the value of Cook's distances to be "large" they should be four times the number of parameters in the model divided by the number of independent samples (i.e. > 0.01, > 0.01, > 0.03, and > 0.07 for FDS, BDS, FCB, and BCB, respectively). Second, data were considered outliers if the value of the absolute DEFETAS exceeding $2/\sqrt{n}$, where n is the number of independent sample (i.e. > 0.07, > 0.08, > 0.11, and > 0.19 for FDS, BDS, FCB, and BCB, respectively). Third, outliers were also defined as when the value of the absolute DFFITS exceeded $2\sqrt{p/n}$, where p is the number of parameter in the model and n is the number of independent sample (i.e. > 0.10, > 0.12, > 0.16, and > 0.27 for FDS, BDS, FCB, and BCB, respectively). Applying these principles, outliers being detected from 30 FDS, 12 BDS, 10 FCB, and 6 BCB studies and data from these studies were then deleted from the final dataset. Following removal of the outliers, some changes in the observed relationship between the mean memory test scores and years of publication were found (r_{FDS} with vs without outlier = 0.12** vs 0.12**, r_{BDS} with vs without outlier = -0.10** vs -0.06**, r_{FCB} with vs without outlier = 0.13** vs 0.10**, and r_{BCB} with vs without outlier = 0.01^{ns} vs -0.17**). However, these numerical changes concerning r_{FDS} , r_{BDS} , r_{FCB} , and r_{BCB} showed no significant differences ($Z_{FDS} \leq 0.01$, $p = 1.00$; $Z_{BDS} = -0.70$, $p = 0.48$; $Z_{FCB} = 0.38$, $p = 0.70$; and $Z_{BCB} = 1.38$, $p = 0.17$) and the effect sizes were small ($d_{FDS} \leq 0.01$, $d_{BDS} = 0.05$, $d_{FCB} = 0.04$, and $d_{BCB} = 0.25$) (see Supplementary information: S1).

In addition, the findings from the current dataset as to publication status (i.e. weighted means for memory tests from published vs. unpublished sources) did not show any significant differences and the magnitudes of difference as indicated by effect sizes (i.e. partial η^2) were small (Cohen, Miles, & Shevlin, 2001) for FDS ($p = 0.15$, partial $\eta^2 < 0.01$), BDS ($p = 0.14$, partial $\eta^2 < 0.01$), and BCB ($p = 0.49$, partial $\eta^2 < 0.01$), respectively. Nevertheless, the weighted mean score for FCB indicated a significant difference between published and unpublished sources but the effect size was also small ($p < 0.05$, partial $\eta^2 = 0.04$).

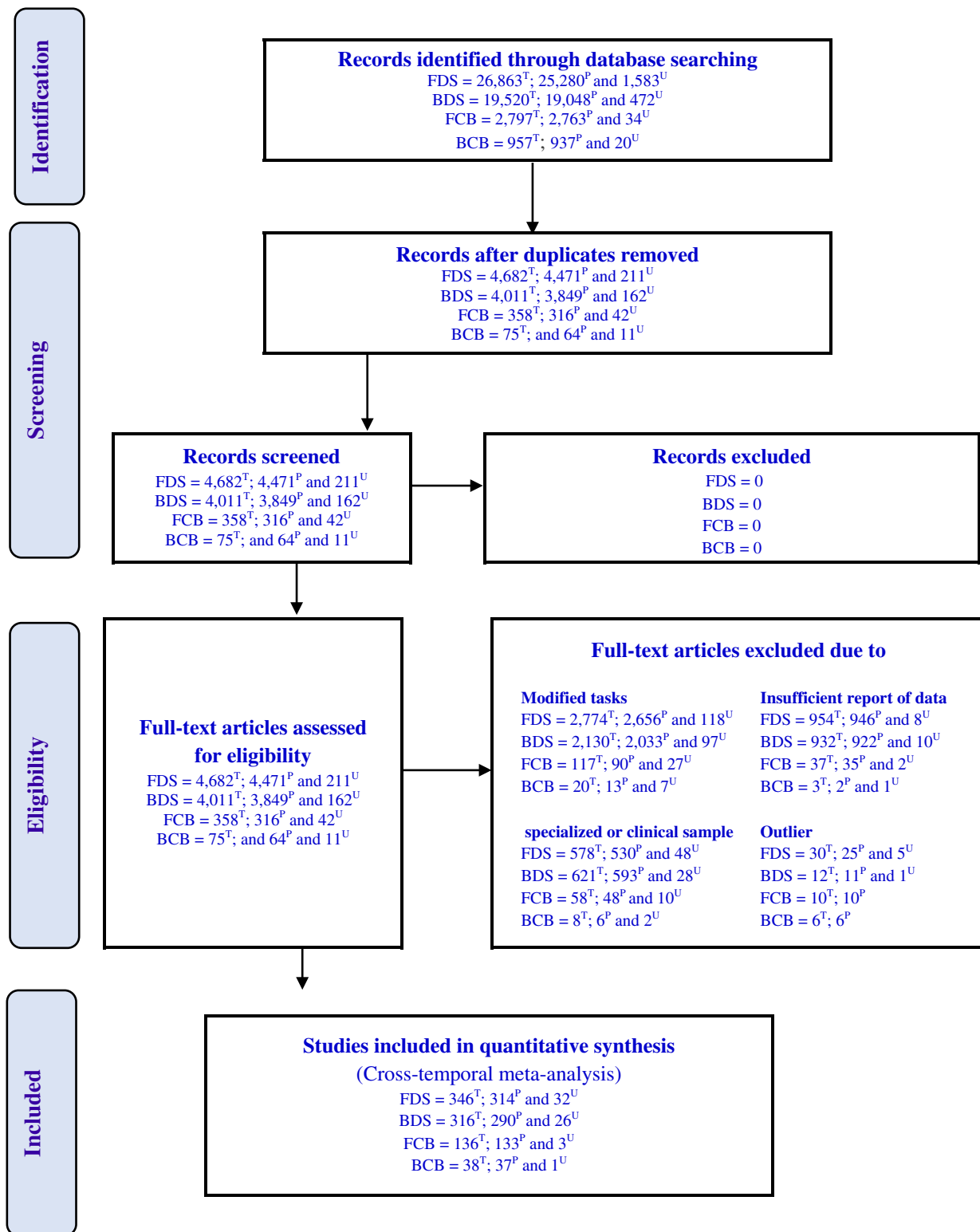


Fig. 1. The PRISMA diagram with literature search results.

Note. FDS = Forward Digit Span, BDS = Backward Digit Span, FCB = Forward Corsi Block, and BCB = Backward Corsi Block. T = Total, P = Published, and U = Unpublished studies.

2.5. Final sample

The literature review processes according to PRISMA guidelines are shown in Fig. 1. The 15-electronic database search yielded 26,863 studies for FDS (94.11% published source), 19,520 studies for BDS (97.58% published source), 2797 studies for FCB (98.78% published source), and 957 studies for BCB (97.91% published source). Of these,

duplicated and unrelated studies were removed after title, keyword, and abstract review. Only, 4682 (95.49% published source), 4011 (95.96% published source), 358 (88.27% published source), and 75 studies (85.33% published source) for FDS, BDS, FCB, and BCB remained for the next step. Subsequently, 346 (90.75% published source), 316 (91.77% published source), 136 (97.79% published source), and 38 (97.37% published source) studies for FDS, BDS, FCB, and BCB

respectively fulfilled the inclusion and exclusion criteria, resulting in 742, 594, 307, and 111 independent samples for FDS, BDS, FCB, and BCB) and 139,677 participants overall (48,955, 70,424, 16,514, and 3784 participants for FDS, BDS, FCB, BCB). References for studies included in this CTMA can be found in the supporting information and the dataset file is openly available at <https://osf.io/4zesc/>.

3. Results

Descriptive statistics concerning years of publication, age of participant, sex, sample size, test platform, and type of country are given in Table 1 for each memory test. The median years of publication were as follows: for FDS, 2003.15 (IQR = 15; span 43 years; 1973–2016), for BDS 2005.33 (IQR = 15; span 41 years, 1975–2016), for FCB, 2008.22 (IQR = 9; span 34 years, 1982–2016), and for BCB 2009.71 (IQR = 6; span 27 years, 1989–2016). The weighted mean ages of the participants were; FDS: 36.76 (SD = 230.90; span 3.65–92.40 years), BDS: 41.27 (SD = 287.11; span 4.00–92.50), FCB: 29.31 (SD = 162.48; span 3.04–86.65); and BCB: 43.06 (SD = 149.39; span 4.00–86.65). The weighted mean scores were; FDS: 6.11 (SD = 7.22); BDS: 4.51 (SD = 9.88); FCB: 5.25 (SD = 6.35); and BCB: 4.74 (SD = 4.81). In addition, the weighted means of SD were; FDS: 1.33 (SD = 3.41); BDS: 1.38 (SD = 4.34); FCB: 1.11 (SD = 3.25); and BCB: 1.01 (SD = 1.90).

Moreover, the weighted mean scores for the four memory tests additionally demonstrated that female participants had higher mean scores than male participants for FDS and BDS (6.12 vs. 5.77 and 4.64 vs. 4.30), but male participants had higher mean scores than female participants for FCB and BCB (4.86 vs. 4.80 and 5.02 vs. 4.75). The average sample sizes for FDS, BDS, FCB and BCB respectively were 65.98 (SD = 191.85), 118.56 (SD = 352.58), 53.79 (SD = 80.52), and 34.09 (SD = 39.56). The most frequent types of test platform for all STM and WM tests were non-computerized versions (for FDS, BDS, FCB and BCB respectively, 91.10%, 92.90%, 74.60%, and 89.20%) and the majority of data for all memory tests were collected from developed countries (respectively 90.70%, 89.10%, 92.20%, and 78.40%).

To test the hypotheses that STM and WM increase or decrease overtime, the correlation coefficients from four weighted simple regression analysis were estimated (reported in Tables 2A and 2B and Fig. 2). It was found that the mean scores for the FDS and FCB measures correlated positively with year of publication (respectively: $r = 0.12$, $p < 0.01$; $r = 0.10$, $p < 0.01$; both for sample size weight). A negative correlation was found between the mean score for BDS and BCS

and year of publication (respectively; $r = -0.06$ for sample size weight, $p < 0.01$; $r = -0.17$ for sample size weight, $p < 0.01$; both for sample size weight).

The analysis tested the moderating effects of the relationship between the mean scores for all memory tests and year of publication, the main finding (see Tables 2A and 2B) being that the moderators were significant predictors for all memory tests. Specifically, the significant correlations between age of participant and FDS ($\beta_{FDS} = 0.14$, $p < 0.01$), BDS ($\beta_{BDS} = 0.17$, $p < 0.01$), and FCB ($\beta_{FCB} = 0.06$, $p < 0.01$) mean scores indicated higher test performances for older participants. Nonetheless, a negative and significant relationship was found between age of participant and the BCB mean score ($\beta_{BCB} = -0.45$, $p < 0.01$), this indicating higher test performance for younger participants. Further, a positive and significant correlation was observed between male percentage and the BCB means score ($\beta_{BCB} = 0.10$, $p < 0.01$), showing higher test performances in male participants on this memory test. However, negative directions were found for the relationships between male percentage and FDS ($\beta_{FDS} = -0.03$, $p < 0.01$), BDS ($\beta_{BDS} = -0.01$, $p < 0.01$), and FCB ($\beta_{FCB} = -0.06$, $p < 0.01$) mean scores, indicating better performance for female participants on these three memory tests.

Significant negative relationships between test platform and FDS ($\beta_{FDS} = -0.02$, $p < 0.01$), BDS ($\beta_{BDS} = -0.03$, $p < 0.01$), FCB ($\beta_{FCB} = -0.08$, $p < 0.01$), and BCB ($\beta_{BCB} = -0.14$, $p < 0.01$) mean scores suggested higher test score of the computerized platform in comparison to the non-computerized platform on all four memory tests. Also, it is clear from the current analysis that positive and significant correlations were found between type of country and FDS ($\beta_{FDS} = 0.16$, $p < 0.01$), BDS ($\beta_{BDS} = 0.26$, $p < 0.01$), FCB ($\beta_{FCB} = 0.07$, $p < 0.01$), and BCB ($\beta_{BCB} = 0.30$, $p < 0.01$) mean scores, indicating better test scores for the participants from developed countries than those from developing countries on the four memory tests. In addition, according to the multiple weighted regression, the relationships between years of publication and mean scores were still statistically significant for all memory tests ($\beta_{FDS} = 0.14$, $p < 0.01$, $\beta_{BDS} < -0.02$, $p < 0.01$, $\beta_{FCB} = 0.08$, $p < 0.01$, and $\beta_{BCB} = -0.03$, $p < 0.01$) after the moderating effects of participant age, male percentage, type of platform, and type of country were taken into account.

4. Discussion

The aim of the study was to determine whether the mean scores on

Table 2A

Single and multiple weighted least squares regression analyses as weighted by sample size for forward digit span and backward digit span scores and score changes per decade.

| Target variables | Ns | b (SE) | β (SE) | t | p | r (R ²) |
|-------------------------|--------|------------------|------------------|----------|--------|---------------------|
| (1) FDS | | | | | | |
| Single regression | 48,955 | | | | | |
| Intercept | | -17.43 (0.85) | | -20.40** | < 0.01 | 0.12 (0.02)** |
| Year of publication | | 0.01 (< 0.01) | 0.12 (< 0.01) | 27.56** | < 0.01 | |
| Multiple regression | 48,995 | | | | | 0.26 (0.07)** |
| Intercept | | -21.16 (0.85) | | -24.89** | < 0.01 | |
| Year of publication (Y) | | 0.01 (< 0.01) | 0.14 (< 0.01) | 31.36** | < 0.01 | |
| Mean age (A) | | < 0.01 (< 0.01) | 0.14 (< 0.01) | 30.91** | < 0.01 | |
| Male percentage (M) | | < -0.01 (< 0.01) | -0.03 (< 0.01) | -6.94** | < 0.01 | |
| Platform (P) | | < -0.05 (< 0.01) | < -0.02 (< 0.01) | -3.88** | < 0.01 | |
| Type of country (TC) | | < 0.67 (0.02) | 0.16 (< 0.01) | 36.97** | < 0.01 | |
| (2) BDS | | | | | | |
| Single regression | 70,424 | | | | | -0.06 (< 0.01)** |
| Intercept | | 16.11 (0.76) | | 21.24** | < 0.01 | |
| Year of publication | | -0.01 (< 0.01) | -0.06 (< 0.01) | -15.29* | < 0.01 | |
| Multiple regression | 70,424 | | | | | 0.32 (0.10)** |
| Intercept | | 7.74 (0.73) | | 10.52 | < 0.01 | |
| Year of publication (Y) | | < -0.01 (< 0.01) | -0.02 (< 0.01) | -5.67** | < 0.01 | |
| Mean age (A) | | < 0.01 (< 0.01) | 0.17 (< 0.01) | 45.20** | < 0.01 | |
| Male percentage (M) | | < -0.02 (< 0.01) | -0.01 (< 0.01) | -4.60** | < 0.01 | |
| Platform (P) | | -0.09 (0.01) | -0.03 (< 0.01) | -7.59** | < 0.01 | |
| Type of country (TC) | | 0.88 (0.01) | 0.26 (0.01) | 73.20** | < 0.01 | |

Table 2B

Single and multiple weighted least squares regression analyses as weighted by sample size for forward Corsi block span and backward Corsi block span scores and score changes per decade.

| Target variables | Ns | b (SE) | β (95% CI) | t | p | r (R ²) |
|-------------------------|--------|------------------|------------------|----------|--------|---------------------|
| (1) FCB | | | | | | |
| Single regression | 16,514 | | | | | |
| Intercept | | -13.47 (1.50) | | -8.98** | < 0.01 | 0.10 (0.01)** |
| Year of publication | | 0.01 (< 0.01) | 0.10 (< 0.01) | 12.48** | < 0.01 | |
| Multiple regression | 16,514 | | | | | 0.18 (0.03)** |
| Intercept | | -11.44 (1.54) | | -7.44** | < 0.01 | |
| Year of publication (Y) | | 0.01 (< 0.01) | 0.08 (0.01) | 10.75** | < 0.01 | |
| Mean age (A) | | < 0.01 (< 0.01) | 0.06 (0.01) | 7.38** | < 0.01 | |
| Male percentage (M) | | < -0.01 (< 0.01) | -0.06 (0.01) | -8.34** | < 0.01 | |
| Platform (P) | | -0.15 (0.01) | -0.08 (0.01) | -10.14** | < 0.01 | |
| Type of country (TC) | | 0.36 (0.04) | 0.07 (0.01) | 9.64** | < 0.01 | |
| (2) BCB | | | | | | |
| Single regression | 3784 | | | | | |
| Intercept | | 46.54 (3.81) | | 12.20 | < 0.01 | -0.17 (0.03)** |
| Year of publication | | -0.02 (< 0.01) | -0.17 (0.02) | -10.96** | < 0.01 | |
| Multiple regression | 3784 | | | | | 0.51 (0.26)** |
| Intercept | | 12.45 (3.56) | -0.03 (0.02) | 3.50 | < 0.01 | |
| Year of publication (Y) | | < -0.01 (< 0.01) | -0.45 (0.02) | -2.13** | < 0.01 | |
| Mean age (A) | | < -0.01 (< 0.01) | 0.10 (0.01) | -28.50** | < 0.01 | |
| Male percentage (M) | | < 0.01 (< 0.01) | -0.14 (0.02) | 7.04* | < 0.01 | |
| Platform (P) | | -0.39 (0.04) | 0.30 (0.01) | -9.51** | < 0.01 | |
| Type of country (TC) | | 0.73 (0.04) | -0.03 (0.02) | 20.16** | < 0.01 | |

Note Ns = number of participants, *p < 0.05, and **p < 0.01.

tests of verbal and visuospatial STM and WM change over decades, in line with the FE and the anti-FE. Accordingly, the FE for STM and anti-FE for WM were investigated using CTMA on 1754 independent samples. The results revealed the FE for the verbal and visuospatial STM measures (i.e. FDS and FCB) and anti-FE for the verbal and visuospatial WM measure (i.e. BDS and BCB). Both verbal and visuospatial STM measures are possibly sensitive to the FE in which the mean scores for FDS and FCB increase over time ($r_s = 0.12$ and 0.10 , $p_s < 0.01$). In contrast, both verbal and visuospatial WM measures are probably sensitive to the anti-FE in that the mean scores for BDS and BCB decrease over time ($r_s = -0.06$ and -0.17 , $p_s < 0.01$). The gain on the verbal and visuospatial STM tasks (i.e. FDS and FCB) and decline on the verbal and visuospatial WM tasks (i.e. BDS and BCB) are consistent with the co-occurrence model (Woodley of Menie & Fernandes, 2015). Based on this model, it is predicted that less g-loaded measures should show a rise, whereas more g-loaded measures may show the opposite trend as shown in the studies by Gignac (2015) and Woodley of Menie and Fernandes (2015).

Some studies have also supported the notion that FDS and BDS are differentially related to g, namely that BDS is essentially a more attentionally demanding task than FDS, such that BDS is more strongly

related to g (Ganzach, 2016; Jensen & Figueroa, 1975). In addition, the less g-loaded indicator implies a major environmental influence, with the FE gain predominantly driven by environmental causes (te Nijenhuis & van der Flier, 2013), with additional practice and training effects (Bartels, Wegrzyn, Wiedler, Ackermann, & Ehrenreich, 2010; Chase & Ericsson, 1981; Roediger & Karpicke, 2006; St Clair-Thompson & Holmes, 2008). Furthermore, the existence of the FE for STM tasks (i.e. FDS and FCB) possibly adds robustness to the overall FE phenomenon, as it demonstrates in this study that the FE occurs on ratio-scale measures, and does not simply emerge as an artefact of comparing between cohorts measured on interval scale measures of cognitive ability (Woodley of Menie, Peñaherrera, Fernandes, Becker, & Flynn, 2016).

In comparison to the FDS task, the performance score on the more g-loaded BDS measure may be potentially related to genetic-driven factors (Ando, Ono, & Wright, 2001; Karlsgodt, Bachman, Winkler, Bearden, & Glahn, 2011; Karlsgodt et al., 2010; Kremen et al., 2007). Galton (1869) was the first to propose the existence of dysgenic effects on cognitive abilities and this has been considered in detail by Lynn (2011), with Woodley of Menie and Fernandes (2015) suggesting that the negative trend for BDS score concerning year of publication could

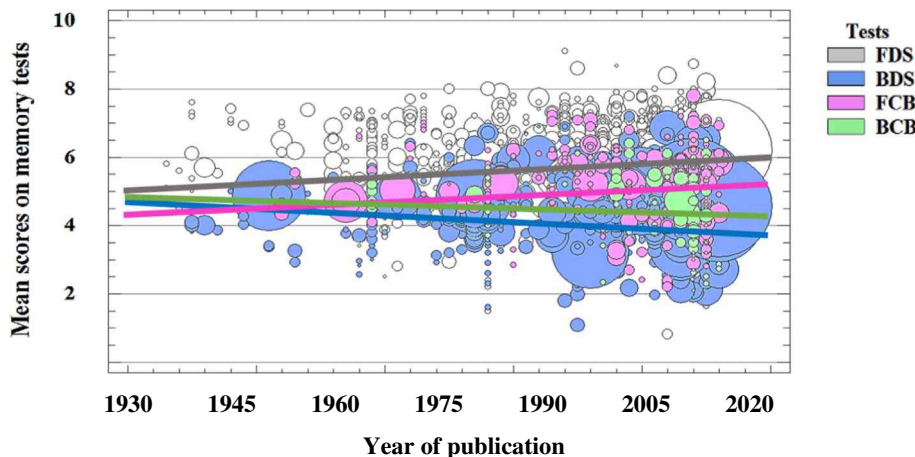


Fig. 2. Weighted simple regression of the relationship between the mean scores of four memory tests (FDS, BDS, FCB, and BCB) and year of publication. The bubble colours represent different types of memory tests, and the sizes provide a visual analogue of the relative sample size. (For interpretation of the references to color in this figure legend, the reader is referred to the web version of this article.)

also be explained by dysgenic fertility on the more g-loaded tasks, as found in several studies (Peach, Lyerly, & Reeve, 2014; Woodley of Menie & Dunkel, 2015; Woodley of Menie, Fernandes, Figueredo, & Meisenberg, 2015; Woodley of Menie, Figueredo, Dunkel, & Madison, 2015; Woodley of Menie & Meisenberg, 2013). That is, it has been proposed that more intelligent men and women in developed countries tend to have fewer offspring and also more highly educated women are more likely to desire childlessness (Kanazawa, 2014; Livingston & Cohn, 2010; Meisenberg, 2010). Speculatively, this might support a steady decline in cognitive ability in some advanced industrial nations (Kanazawa, 2014). Using data from a handful of developed and developing countries, several attempts have been made to estimate the theoretical IQ lost due to the action of dysgenic selection, these countries including the UK and US (for a meta-analysis see Woodley of Menie, 2015), Taiwan (Chen, Chen, Liao, & Chen, 2013), and Kuwait (Abdel-Khalek & Lynn, 2008). Furthermore, anti-FEs on pencil-and-paper tests have also been observed across several different ability measures and in several countries (for reviews, see Dutton & Lynn, 2013, 2015). These IQ declines, like dysgenic effects, tend to be biggest when estimated using more g loaded subtests (Woodley & Meisenberg, 2013; Woodley of Menie & Dunkel, 2015).

Moreover, there is also evidence from the meta-analysis that the WM span is significantly more age sensitive than the STM span (Bopp & Verhaeghen, 2005), suggesting that WM is more susceptible to decline than STM with advanced age, confirming previous suggestions (Bowles & Salthouse, 2003; Hale et al., 2011; McNab et al., 2015; Salthouse, 1994; Wingfield et al., 1988). The increasing longevity in developed countries has resulted in a higher proportion of aging population (Ediev, 2011). For instance, mortality at age 75 years or more in the UK accounted for approximately 12% of all deaths of the last century; and it was around 39% in 1951, dramatically rising in the 21st Century to 65% in 2004 (Howse, 2006). In the current dataset, the samples were composed of 29.9% from '> 60 years old', followed by 21.2% from '13–25 years old', and 19.6% from '7–12 years old', respectively. Based on a lifespan perspective of age-related cognitive decline, it is plausible that WM decline over time may stem in part from the performance scores from this older age group across the more economically developed nations.

In addition, with respect to the negative relationship between visuospatial WM, as indexed by BCB, and year of publication, this negative change also supports the co-occurrence model (Woodley of Menie & Fernandes, 2015). Visuospatial WM has been found to be a strong predictor of mathematical ability (e.g. Ashkenazi et al., 2013; De Smedt et al., 2009; Hubber, Gilmore, & Cragg, 2014; van der Ven, van der Maas, Straatemeier, & Jansen, 2013). The mathematic scores on the International Mathematical Olympiad (IMO) during 1959–2015 indicate a significant and negative trend between the year of competition and efficiency score (ratio of attained score and all possible score) ($n = 54$, $r = -0.71$, $p < 0.01$; weighted by sample size for each year) (IMO, 2016). Furthermore, the current findings are consistent with many studies that have observed decline on attentionally demand visuospatial tasks (Pietschnig & Gittler, 2015; Woodley of Menie & Fernandes, 2016). Also, it is plausible that a higher proportion of the aging population will play some part in the given negative relationship since it was found that performance on the visual-spatial tasks that recruited attentional resource decreased at faster rates than the verbal tasks as a function of age (Hale et al., 2011). Due to the fact that this is the first study that investigated the FE gain for visuospatial STM and the FE decline for visuospatial WM, further investigations are strongly suggested.

For moderating effects on mean scores of memory tests, the positive relationship between mean age and the mean scores on FDS, BDS, and FCB suggested higher performance of older than younger participants for verbal and visuospatial memory tests. However, the magnitudes of correlation, whilst significant, were small ($r_{FDS} = 0.14$, $r_{BDS} = 0.17$, and $r_{FCB} = 0.06$, all $p < 0.01$). This finding is partially contrary to

prevailing notions as to aging resulting in memory decline (León, Tascón, & Cimadevilla, 2016; Salthouse, 2009). Due to the fact that our dataset contained a high proportion of older participants and they are nearly 90% from developed countries, it is plausible that enriched environment, education, and health care systems in developed countries protected the participants in part from short-term or working memory decline due to aging processes (Fuchs et al., 2016; Leal-Galicia, Castañeda-Bueno, Quiroz-Baez, & Arias, 2008; Nouchi & Kawashima, 2014; Springer, McIntosh, Winocur, & Grady, 2005; Uchida & Kawashima, 2008). On the other hand, a negative effect of aging on memory decline was found in BCB and the magnitude of relationship is moderate ($r_{BCB} = -0.45$, $p < 0.01$). This result is consistent with previous studies that found a more pronounced aging effect on visuospatial than verbal WM (Brockmole & Logie, 2013; Peich, Husain, & Bays, 2013).

Sex differences on STM and WM scores have been demonstrated in the current study. Although higher performances for female participants were observed on FDS, BDS, and FCB, the magnitudes of relationship were marginal ($r_{FDS} = -0.03$, $r_{BDS} = -0.01$, and $r_{FCB} = -0.06$, all $ps < 0.01$). Nevertheless, this finding is supported by previous research indicating women outperformed men on verbal (Lowe et al., 2003) and visual memory measures (Andreano & Cahill, 2009). In fact, with a higher number of older adults in the current dataset, the findings are consistent with a previous study demonstrating older men doing more poorly on memory tests and having smaller hippocampal volumes than older women (DeCarli, 2015), although this might not apply to the current data in that hippocampus is thought to be mainly involved in long-term memory. In contrast, the findings for the BCB were in line with several previous studies suggesting that male participants were superior to women participants (Brunetti, Del Gatto, & Delogu, 2014b; Jones & Healy, 2006). However, this finding may also need to be interpreted with caution because of small independent samples of men and women (13 vs. 15) on BCB.

The study also found that there were higher memory scores on STM and WM tests that use computerized platforms in comparison to 'paper and pencil' versions. As Noyes and Garland (2008) pointed out, many paper-based or non-computerized tasks have been transferred onto computers but there may not be total equivalence. There are few comparison studies, but a better FDS score has been observed for a computer-based version in comparison to the non-computer version (Tractenberg & Freas, 2007). A converse higher numerical FCB score was also shown for the most often used wooden platform but the difference did not achieve statistical significance (Robinson & Brewer, 2016). An important point should be considered when considering the current finding is that generalization may not be warranted because of the disproportionately few independent samples using computerized platforms in our dataset, as shown in Table 1 ($FDS_{computerized} = 8.90\%$, $BDS_{computerized} = 7.10\%$, $FCB_{computerized} = 25.40\%$, and $BCB_{computerized} = 10.80\%$).

The study made a comparison between developed and developing countries, finding that people in developed countries showed higher performance on all four tests. To date, although there has been no previous meta-analysis investigating country type for working memory, a recent study by our group has indicated that participants in developed countries have significantly higher scores on general cognitive ability tests (Wongupparaj et al., 2015). Health and nutrition issues are the main concerns in many developing countries (Müller & Krawinkel, 2005) and these problems have been found to affect the memory span of children (Jukes, 2005; Miu et al., 2016; Niehaus et al., 2002). Furthermore, many studies on illiterate and unschooled people have demonstrated that low scores on STM and WM tasks were usually observed in illiterate people (Kosmidis, Zafiri, & Politimou, 2011). In all, once all moderators (age, sex, test platform and type of country) were included in the model, the relationships between year of publication and mean scores for STM and WM tasks remain statistically significant. This indicated the robustness of the FE and anti-FE on memory measures.

Finally, even with significant FE and anti-FE for verbal/visuospatial STM and for verbal/visuospatial WM in the present study, the FE and anti-FE are relatively small ($r_s = 0.12$ for FDS, 0.10 for FCB, -0.06 for BDS, and -0.17 for BCB, respectively). Only small shifts in verbal STM and WM task performance were also found in Woodley of Menie and Fernandes (2015). Indeed, it may also be the case that the small effect may in part be attributable to the limited range of variability in test scores for the memory tasks considered. For instance, around 90% of healthy adults can recall on DS tasks somewhere between 5 and 8 digits (Wechsler, 1939). In addition, it is also plausible that other underlying reasons may account for the decline of the mean scores for BDS over time, for example, diminishing returns of IQ-boosting effects, global ceiling effects and then reversal of scores on IQ tests (Pietschnig, 2016; Pietschnig & Gittler, 2015).

5. Conclusions

In summary, verbal and visuospatial STM shows a gradual rise, supporting the FE for STM as measured by FDS and FCB, whilst verbal and visuospatial WM gradually declines over the past four decades, supporting the anti-FE for WM as indexed by BDS and BCB. Over time, environmental influences might have driven the test score changes for both types of STM but speculatively, dysgenic selection against general intelligence and also possibly age-related cognitive decline could have influenced the declining test score for verbal and visuospatial WM, especially for developed countries. These patterns of the results are in line with the predictions from the co-occurrence model, that is, the FE effect possibly occurs on less g-loaded abilities, whereas the anti-FE effect may concentrate on high g-loaded abilities.

However, these analyses only examined performances on digit span and Corsi block task to measure respectively verbal and visuospatial STM and WM. Further research should explore other standard STM and WM measures to explore and to extend the current findings. For example, this might include word span (Baddeley, 1986; Tallnd, 1965) and letter span (Kinsbourne, 1974; Taub, 1975) for verbal STM, the visual pattern test (Della Sala, Gray, Baddeley, Allamano, & Wilson, 1999) for spatial STM, reading span (Daneman & Carpenter, 1980) for verbal WM, and n-back tasks (Jaeggi, Buschkuhl, Perrig, & Meier, 2010; Kirchner, 1958) for verbal and visuospatial WM. Such studies would help to further elucidate the potential FE mechanisms of change in relation to STM and WM.

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Conflict of interest

The authors declare no conflict of interest.

Contributors

Peera Wongupparaj, Veena Kumari and Robin G. Morris designed the study. Peera and Rangirat Wongupparaj undertook the statistical analysis, data collection and prepared the first draft. All authors contributed to and approved the final manuscript.

Appendix A. Supplementary data

Supplementary data to this article can be found online at <http://dx.doi.org/10.1016/j.intell.2017.07.006>.

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